

Estimating Poverty Measures using Truncated Distributions and its Applications

Nagwa Albehery, Tonghui Wang

Abstract- Poverty measures are used to measure poverty levels or degrees of poverty in a population. In this paper, we investigate the estimation of poverty measures using truncated log-normal, truncated gamma, and truncated epsilon-skew-normal distributions. For comparisons and illustrations, our results are applied to the real data sets collected in Egypt between 1995/1996 and 2008/2009.

Index Terms: Poverty measures, Parametric estimation, Income distributions, Truncated distributions.

I. INTRODUCTION

Poverty measure is defined by Duclos (2006) as the function of a given distribution $F(x)$ of random variable X (income or expenditure) and a poverty line z , denoted by $P(F, z)$ where $x < z$, and the poverty line z is a borderline between poor and non-poor people in a population. There are two common classes of poverty measures.

(a) Additively separable poverty measures. The class of additively separable poverty measures are defined by a given poverty line z and an individual deprivation function $h(x, z)$ that is differentiable in both x and z for $x \in [0, z]$ and $z \in (0, +\infty)$. This class of poverty measures is decomposable across population subgroups. The indices of this class have the property of being expressible as a weighted sum (as separable function) and have the following form

$$P(F, z) = \int_0^{+\infty} I_{[0,z]}(x) h(x, z) dF(x), \quad (1)$$

where the indicator function $I_{[0,z]}(x)$ is defined by

$$I_{[0,z]}(x) = \begin{cases} 1 & x \in [0, z] \\ 0 & \text{elsewhere.} \end{cases}$$

The most popular index in this class is called Foster's index (Foster, 1984), in which $h(x, z)$ is given by
Thus the Foster's index is defined as

$$h(x, z) = \left(1 - \frac{x}{z}\right)^\alpha, \quad \alpha \geq 0.$$

When $\alpha = 0$, the Foster's index is called the Head-count ratio, denoted by H . When $\alpha = 1$, it is called the

Poverty gap ratio, denoted by $P_1(F, z)$. When $\alpha = 2$, it is called the Severity of poverty, denoted by $P_2(F, z)$.

(b) Rank-based poverty measures. Rank-based Poverty measures use the position or the rank of each poor as an indicator of the relative deprivation function $q(x, z)$ that is differentiable in both x and z for $x \in [0, z]$ and $z \in (0, \infty)$, and has the following form

$$P(F, z) = \int_0^{+\infty} I_{[0,z]}(x) q(x, z) dF(x). \quad (3)$$

As an example, Sen (1976) proposed the following relative deprivation function

$$q(x, z) = 2 \left(\frac{z - x}{z} \right) \left(1 - \frac{F(x)}{F(z)} \right),$$

and the index, called Sen's index later, is

$$S = 2 \int_0^{+\infty} I_{[0,z]}(x) \left(\frac{z - x}{z} \right) \left(1 - \frac{F(x)}{F(z)} \right) dF(x). \quad (4)$$

In addition, the Modified Sen's index (Shorrocks, 1995) is defined by

$$S_m = 2 \int_0^{+\infty} I_{[0,z]}(x) \left(\frac{z - x}{z} \right) (1 - F(x)) dF(x). \quad (5)$$

In this paper, a brief review of income distribution models is given in Section 2. Truncated distributions and parameters estimation are discussed in Section 3. Truncated distributions to estimate poverty measures are investigated in Section 4. In Section 5, our main results are applied to the real survey data on household expenditures collected in Egypt (from Egyptian Central Agency of Statistics) between 1995/1996 and 2008/2009, as an illustration.

$$g_X(x) = \begin{cases} \frac{A^{-1}}{x} \exp\left(-\frac{(\ln x - \mu)^2}{2\sigma^2}\right) & \text{if } 0 < x \leq T \\ 0 & \text{elsewhere,} \end{cases}$$

II. A BRIEF REVIEW OF INCOME DISTRIBUTION MODELS

As mentioned before, an estimation of income distributions is required to estimate poverty measures. Two approaches are used to estimate the income distribution.

(a) The first approach. This approach is a non-parametric approach which is easily understood and allows data to explain itself using histograms. The empirical distribution function is defined by

$$\hat{F}_n(t) = \frac{1}{n} \sum_{i=1}^n I_{[0,t]}(X_i),$$

Manuscript Received on April 2015.

Nagwa Albehery, Department of mathematics and Applied Statistics and Insurance, Helwan University/ Faculty of Commerce and Business administration, Cairo, Egypt.

Tonghui Wang, Department of Mathematics, New Mexico State University/ College of Sciences/ Las Cruces, United States of America.

the most important non-parametric methods that is used to estimate the cumulative distribution function $F(x)$ of random variable X (income or expenditure) and it is used to estimate the poverty measures which is defined in (1) or in (3), (see Kakwani (1993), Davidson (2000) and Berger (2003) for details). Also, the kernel method is a good tool in non-parametric approach to estimate the distribution of X and is defined by

$$\hat{f}(x) = \frac{1}{nh} \sum_{i=1}^n \psi\left(\frac{x - x_i}{h}\right),$$

Where h is the window width and ψ is the kernel function which is usually, but not always, symmetric, see Zheng (2001) and Berger (2003).

(b) The second approach. This approach is a parametric approach which uses the parametric models of income distributions. These models assume that the income distribution follows a known functional form with unknown parameters. The pareto, log-normal, beta and gamma distributions are the most popular examples of the distribution of income, see Singh (1976), Salem (1974), Harrison (1981), and MacDonald (1984, 1995). In the literature examined, many distributions have been used to describe the distribution of income. In 1897, the distribution of income in the form of a probability density function was proposed by Vilfredo Pareto, see Harrison (1981). Pareto's function is accurately fitting high levels of income, but it is not good in describing the law end of the distribution. The log-normal distribution fits the lower levels of income better than the upper levels of income, see Singh (1976) and Harrison (1981). The log-normal density function takes the following form

$$f_X(x) = \begin{cases} \frac{1}{x\sigma\sqrt{2\pi}} \exp\left(-\frac{(\ln x - \mu)^2}{2\sigma^2}\right) & \text{if } x > 0 \\ 0 & \text{elsewhere,} \end{cases}$$

Where $\mu \in \mathbb{R}$ and $\sigma > 0$ are the mean and the variance of the normal distribution. Salem (1974) introduced the gamma distribution to fit United States income data. The gamma density function may take this form

$$f_X(x) = \begin{cases} \frac{\beta^\alpha}{\Gamma(\alpha)} (x - c)^{\alpha-1} e^{-\beta(x-c)} & \text{if } x \geq c \\ 0 & \text{elsewhere,} \end{cases}$$

Where α and β are two positive parameters and $c > 0$ is a constant. In the next section truncated distributions and parameters estimation are discussed.

III. TRUNCATED DISTRIBUTIONS AND PARAMETERS ESTIMATION

The truncated distributions have many applications in science. In this section truncated log-normal, truncated gamma, and truncated epsilon-skew-normal distributions are defined. Also, the parameters estimation of these distributions are discussed.

(i) Truncated log-normal distribution. The truncated log-normal distribution has a density function which may be written by

$$g_X(x) = \begin{cases} \frac{A^{-1}}{x} \exp\left(-\frac{(\ln x - \mu)^2}{2\sigma^2}\right) & \text{if } 0 < x \leq T \\ 0 & \text{elsewhere,} \end{cases}$$

Where

$$A = \int_0^T \exp\left(-\frac{(\ln x - \mu)^2}{2\sigma^2}\right) dx.$$

(ii) Truncated gamma distribution. The truncated gamma density function may take this form

$$g_X(x) = \begin{cases} A^{-1} (x - c)^{\alpha-1} e^{-\beta(x-c)} & \text{if } c \leq x \leq T \\ 0 & \text{elsewhere,} \end{cases}$$

Where $c > 0$ and

$$A = \int_c^T (x - c)^{\alpha-1} e^{-\beta(x-c)} dx.$$

(iii) Truncated epsilon-skew-normal distribution. The epsilon-skew-normal distribution proposed by Mudholkar (2000) has a probability density function written by

$$f_X(x) = \begin{cases} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2(1-\varepsilon)^2}\right) & \text{if } x < 0 \\ \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2(1+\varepsilon)^2}\right) & \text{if } x \geq 0 \end{cases}$$

Where $\theta=0$ and $\sigma=1$. Then the truncated epsilon-skew-normal probability distribution can be written by

$$g_X(x) = \begin{cases} A^{-1} \exp\left(-\frac{x^2}{2(1+\varepsilon)^2}\right) & \text{if } 0 \leq x \leq T \\ 0 & \text{elsewhere,} \end{cases}$$

Where

$$A = \int_0^T \exp\left(-\frac{x^2}{2(1-\varepsilon)^2}\right) dx.$$

(a) Estimating the parameters of truncated distributions.

In the literature reviewed, least-squares procedure and maximum likelihood estimation are two common methods that are used to estimate the parameters of statistical models and truncated distributions, see Chapman (1965), Salem (1974), Singh (1976), Slocomb (1977), Harrison (1981), Bandourian (2000), Mudholkar (2000), and Aban (2006). In this section, the parameters of truncated log-normal, truncated gamma, and truncated epsilon-skew-normal distributions are estimated.

(1) Method of least-squares procedure. Parameters of the truncated distributions are estimated by minimizing the mean squared error fit on a plot of the truncated distribution function see Chapman (1965) and Aban (2006). Let n_i be the observations be grouped by classes

$$(a_i - h_i, a_i + h_i), i = 1, \dots, r,$$

Where

$$a_1 - h_1 = 0, a_r + h_r = T, a_i + h_i = a_{i+1} - h_{i+1},$$

Let $n_{(i)}$ is the number of observations falling in class i

between $a_{(i)} - h_{(i)}$ and $a_{(i)} + h_{(i)}$ and

$$n = \sum_{i=1}^r n_i.$$

Define

$$p_i = \int_{a_i-h_i}^{a_i+h_i} g_X(x)dx \doteq g(a_i)(2h_i),$$

and

$$q_i = \frac{n_i}{n}.$$

Now, we can use the form

$$\ln p_i - \ln p_{i+1}, \quad i = 1, \dots, r-1$$

to estimate the parameters of the truncated distributions with replacing $q_{(i)}$ instead of $p_{(i)}$, $i=1, \dots, r-1$. The least-squares procedure is a common and a good method that is used to estimate the parameters of the truncated distributions when the income or expenditure data are grouped in intervals.

(2) Method of maximum likelihood estimation. For a fixed data set and a given statistical model, the maximum likelihood provides estimates for the model's parameters that maximize the likelihood function $L(x, \theta)$ defined by

$$L(x, \theta) = \prod_{i=1}^r g(x_i, \theta),$$

Where θ is the parameter and $g(x_i, \theta)$ is the probability densities function of the truncated distribution.

IV. TRUNCATED DISTRIBUTIONS TO ESTIMATE POVERTY MEASURES

As mentioned before, a poverty measure is defined as the function of a given distribution $F(x)$ of random variable X (income or expenditure) and a poverty line z . We can conclude that the poverty measures depend on the low levels of income or expenditure; the left part of the distribution $f(x)$, $x < z$. The truncated log-normal, truncated gamma, and truncated epsilon-skew-normal distributions then can be investigated to describe the distribution of the poor in a population and to estimate the poverty measures. The sum of squared errors (SSE) (Bandourian (2000)), the sum of absolute errors (SAE) (Chang (2008)), and Chi-square goodness of fit (Harrison (1981)) are three measures; that are used to evaluate the performance of the estimated truncated distributions and defined by

$$SSE = \sum_{i=1}^r (q_i - p_i(\hat{\theta}))^2, \quad SAE = \sum_{i=1}^r |q_i - p_i(\hat{\theta})|,$$

and

$$\chi^2 = \sum_{i=1}^r \left[\frac{(q_i - p_i(\hat{\theta}))^2}{p_i(\hat{\theta})} \right],$$

Where r is the number of income or expenditure's groups and $\hat{\theta}$ is the estimated parameters vector of the truncated distribution. The selected truncated distribution will be used to estimate the Foster poverty indices, Sen index, and modified Sen index given in (2), (4), and (5), respectively.

V. REAL DATA APPLICATIONS

The most important data sources to measure the poverty levels in Egypt are household surveys. Egyptian household surveys are available for the years 1995/1996, 1999/2000, 2004/2005 and 2008/2009. Each period is written as two years because each survey starts in mid-year of the first period and ends in mid-year of the second one. The title of the surveys for 1995/1996, 1999/2000, 2004/2005 and 2008/2009 was "The Research of Consumption and Expenditure in Egypt" and all households surveys are produced by the Egyptian Central Agency of Statistics (ECAS). In this paper, we will use household expenditure instead of income because many reported incomes in developing countries might be far less than real incomes. Also, in the developing countries, the income data is limited since many people don't report secondary sources of income. In our work, consumption expenditure data sets on urban areas and on rural areas are collected independently. Also for each consumption expenditure group the average expenditure \bar{X} , the average expenditure on food \bar{X}_f , the number of households NH, and the number of individuals NI are given. The relative poverty line defined by a percentage of mean or median income or expenditure is used as the borderline between the poor and non-poor people see Oti (1990), Gustafsson (1996), Sahn (2000) and Zheng (2001). In this paper, The relative poverty line is estimated as

(a) 1/3 of the median of annual household expenditure, $z_{\{-}}$,

which serves as the minimum poverty line and

(b) 2/3 of the median of annual household expenditure, $z_{\{+}}$,

which serves as the maximum of poverty line. The relative poverty line is updated automatically over time when the expenditure changes. The poverty line for rural, urban, and total Egypt survey data between 1995/1996 and 2008/2009 are estimated and listed in Table 1. The parameters of the truncated gamma, truncated log-normal, and truncated epsilon-skew-normal distributions are estimated using the real data sets collected in Egypt from 1995/1996 to 2008/2009 and using the method of least-squares procedure. Based on the sum of squared errors (SSE), the sum of absolute errors (SAE), and chi-square goodness of fit measures, the truncated gamma distribution has the best performance of the data and will be used to describe the expenditure of the poor, see Figure 1, 2, 3, 4 and Figure 5. The estimated parameters of truncated gamma distribution, the sum of absolute errors (SAE), the sum of squared errors (SSE), and chi-square goodness of fit values are listed in Table 2. Also, the estimated poverty measures using the truncated gamma distribution are listed in Table 3 and Table 4.

Figures and Tables

TABLE I
ESTIMATED RELATIVE POVERTY LINES USING RURAL, URBAN AND TOTAL SURVEY DATA BETWEEN 95/96 AND 08/09.

| Area | 1995/1996 | | 1999/2000 | |
|-------|-----------------------|-----------------------|-----------------------|-----------------------|
| | Relative Poverty Line | Relative Poverty Line | Relative Poverty Line | Relative Poverty Line |
| | z_{-} | z_{+} | z_{-} | z_{+} |
| Rural | 1601.4 | 3202.9 | 2162.5 | 4325.0 |
| Urban | 2061.2 | 4122.3 | 2990.1 | 5980.2 |
| Area | 2004/2005 | | 2008/2009 | |
| | Relative Poverty Line | Relative Poverty Line | Relative Poverty Line | Relative Poverty Line |
| | z_{-} | z_{+} | z_{-} | z_{+} |
| Rural | 2664.6 | 5329.2 | 4371.3 | 8742.6 |
| Urban | 3577.7 | 7155.4 | 5325.0 | 10650.1 |
| Total | 2997.8 | 5995.6 | 4661.8 | 9323.6 |

Estimating Poverty Measures using Truncated Distributions and its Applications

TABLE II

APPROXIMATED PARAMETERS OF TRUNCATED GAMMA, THE SUM ABSOLUTE ERRORS, THE SUM SQUARE ERRORS, AND CHI-SQUARE VALUES USING RURAL, URBAN, AND TOTAL SURVEY DATA BETWEEN 95/96 AND 08/09.

| area | 1995/1996 | | | | |
|-------|-----------|----------|--------|--------|----------|
| | β | α | ASE | SSE | χ^2 |
| Rural | 0.00020 | 2.763 | 0.3648 | 0.0312 | 0.1031 |
| Urban | 0.00030 | 3.458 | 0.2540 | 0.0249 | 0.0711 |
| area | 1999/2000 | | | | |
| | β | α | ASE | SSE | χ^2 |
| Rural | 0.00044 | 3.821 | 0.1460 | 0.0047 | 0.0289 |
| Urban | 0.00060 | 5.029 | 0.1586 | 0.0040 | 0.0303 |
| area | 2004/2005 | | | | |
| | β | α | ASE | SSE | χ^2 |
| Rural | 0.00010 | 2.995 | 0.0762 | 0.0012 | 0.0066 |
| Urban | 0.00090 | 7.534 | 0.2395 | 0.0093 | 0.0751 |
| Total | 0.00013 | 3.118 | 0.0800 | 0.0014 | 0.0088 |
| area | 2008/2009 | | | | |
| | β | α | ASE | SSE | χ^2 |
| Rural | 0.00005 | 2.620 | 0.5148 | 0.0590 | 0.2373 |
| Urban | 0.00002 | 2.846 | 0.2259 | 0.1075 | 0.0619 |
| Total | 0.00004 | 2.829 | 0.1802 | 0.0107 | 0.0454 |

TABLE III

ESTIMATED POVERTY MEASURES USING TRUNCATED GAMMA, RELATIVE POVERTY LINES, RURAL AND URBAN SURVEY DATA BETWEEN 95/96 AND 08/09.

| Measures | 1995/1996 | | | | 1999/2000 | | | |
|----------------|-----------|--------|-------|--------|-----------|--------|-------|--------|
| | Rural | | Urban | | Rural | | Urban | |
| | z- | z+ | z- | z+ | z- | z+ | z- | z+ |
| H | 4.35% | 23.47% | 2.99% | 20.74% | 3.17% | 21.94% | 3.16% | 25.91% |
| P ₁ | 1.20% | 6.82% | 0.72% | 5.50% | 0.74% | 5.88% | 0.64% | 6.69% |
| P ₂ | 0.53% | 3.07% | 0.28% | 2.29% | 0.28% | 2.43% | 0.22% | 3.93% |
| S | 1.81% | 10.07% | 1.16% | 8.48% | 1.20% | 9.05% | 1.11% | 10.64% |
| S _m | 2.36% | 12.19% | 1.43% | 10.03% | 1.46% | 10.62% | 1.27% | 12.11% |
| Measures | 2004/2005 | | | | 2008/2009 | | | |
| | Rural | | Urban | | Rural | | Urban | |
| | z- | z+ | z- | z+ | z- | z+ | z- | z+ |
| H | 2.50% | 16.35% | 1.92% | 26.24% | 4.58% | 24.07% | 2.59% | 17.21% |
| P ₁ | 0.64% | 4.39% | 0.30% | 6.02% | 1.28% | 7.06% | 0.68% | 4.57% |
| P ₂ | 0.27% | 1.87% | 0.08% | 2.10% | 0.58% | 3.21% | 0.28% | 1.94% |
| S | 0.99% | 6.71% | 0.57% | 10.40% | 1.91% | 10.37% | 1.05% | 7.01% |
| S _m | 1.27% | 8.03% | 0.60% | 11.25% | 2.51% | 12.61% | 1.34% | 8.41% |

TABLE IV

ESTIMATED POVERTY MEASURES USING TRUNCATED GAMMA, RELATIVE POVERTY LINES, AND TOTAL SURVEY DATA IN 04/05 AND 08/09.

| Measures | 2004/2005 | | 2008/2009 | |
|----------------|-----------|--------|-----------|--------|
| | Total | | Total | |
| | z- | z+ | z- | z+ |
| H | 2.67% | 17.39% | 2.98% | 18.47% |
| P ₁ | 0.68% | 4.69% | 0.79% | 5.05% |
| P ₂ | 0.28% | 1.99% | 0.34% | 2.19% |
| S | 1.06% | 7.16% | 1.20% | 7.49% |
| S _m | 1.34% | 8.55% | 1.56% | 9.15% |

| z | Relative frequencies *10 ⁵ | Truncated gamma *10 ⁵ |
|------|---------------------------------------|----------------------------------|
| 500 | 0.23 | 0.29 |
| 1500 | 3.35 | 4.11 |
| 2500 | 7.84 | 11.18 |
| 3500 | 13.64 | 18.6 |
| 4500 | 21.97 | 24.34 |
| 5500 | 26.55 | 27.61 |
| 6500 | 26.43 | 28.49 |

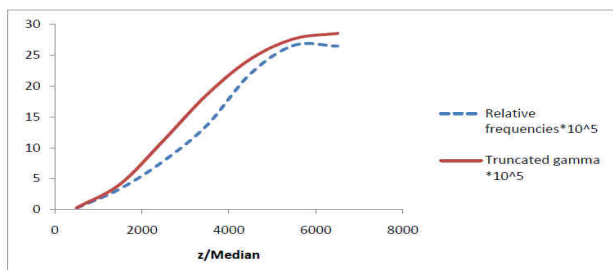


Fig. 1. Truncated gamma distribution - Rural 99/00.

This figure displays the truncated gamma distribution for rural data collected in 1999/2000. It shows that the truncated gamma distribution has the best performance of the data comparing with the relative frequency distribution.

| z | Relative frequencies*10 ⁵ | Truncated gamma *10 ⁵ |
|------|--------------------------------------|----------------------------------|
| 1000 | 0.455 | 0.962 |
| 2500 | 4.337 | 5.15 |
| 3500 | 8.603 | 9.12 |
| 4500 | 14.432 | 13.62 |
| 5500 | 20.455 | 18.39 |
| 6500 | 24.467 | 23.22 |
| 7500 | 25.889 | 27.95 |

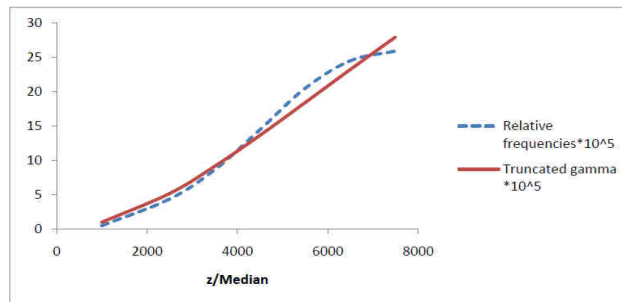


Fig. 2. Truncated gamma distribution - Rural 04/05.

This figure displays the truncated gamma distribution for rural data collected in 2004/2005. It shows that the truncated gamma distribution has the best performance of the data comparing with the relative frequency distribution.

| z | Relative frequencies*10 ⁵ | Truncated gamma *10 ⁵ |
|------|--------------------------------------|----------------------------------|
| 500 | 0.025 | 0.022 |
| 1500 | 1.35 | 1.03 |
| 2500 | 3.57 | 4.43 |
| 3500 | 6.44 | 9.42 |
| 4500 | 10.84 | 14.24 |
| 5500 | 14.72 | 17.54 |
| 6500 | 16.28 | 18.87 |
| 7500 | 16.75 | 18.43 |
| 9000 | 7.51 | 15.62 |

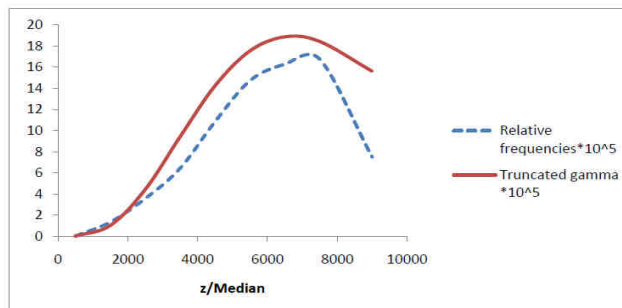


Fig. 3. Truncated gamma distribution - Urban 99/00.

This figure displays the truncated gamma distribution for urban data collected in 1999/2000. It shows that the truncated gamma distribution has the best performance of the data comparing with the relative frequency distribution.

| z | Relative frequencies*10 ⁵ | Truncated gamma *10 ⁵ |
|-------|--------------------------------------|----------------------------------|
| 1000 | 0.155 | 0.012 |
| 2500 | 1.72 | 1.19 |
| 3500 | 3.85 | 4.38 |
| 4500 | 6 | 9.19 |
| 5500 | 9.67 | 13.87 |
| 6500 | 12.4 | 16.79 |
| 7500 | 14.17 | 17.39 |
| 8500 | 15.94 | 16.02 |
| 9500 | 15.23 | 13.47 |
| 10750 | 9.05 | 9.81 |

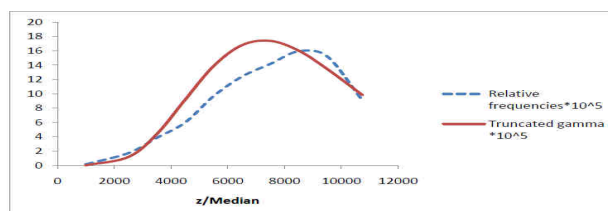


Fig. 4. Truncated gamma distribution - Urban 04/05.

This figure displays the truncated gamma distribution for urban data collected in 2004/2005. It shows that the truncated gamma distribution has the best performance of the data comparing with the relative frequency distribution.

| z | Relative frequencies*10 ⁵ | Truncated gamma *10 ⁵ |
|------|--------------------------------------|----------------------------------|
| 1000 | 0.32 | 0.69 |
| 2500 | 3.2 | 3.92 |
| 3500 | 6.58 | 7.03 |
| 4500 | 10.79 | 10.51 |
| 5500 | 15.92 | 14.11 |
| 6500 | 19.51 | 17.65 |
| 7500 | 21.22 | 20.98 |
| 8500 | 21.47 | 24.01 |

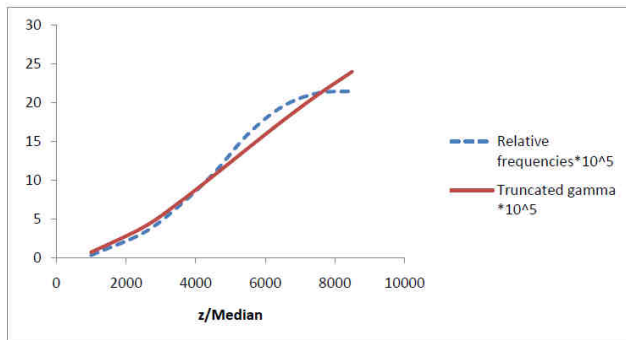


Fig. 5. Truncated gamma distribution - Total 04/05

This figure displays the truncated gamma distribution for total data collected in 2004/2005. It shows that the truncated gamma distribution has the best performance of the data comparing with the relative frequency distribution.

VI. CONCLUSION

The parameters of the truncated gamma, truncated log-normal, and truncated epsilon-skew-normal distributions are estimated using the real data sets collected in Egypt from 1995/1996 to 2008/2009 and using the method of least-squares procedure. Based on the sum of squared errors (SSE), the sum of absolute errors (SAE), and chi-square goodness of fit measures, the truncated gamma distribution has the best performance of the data and will be used to describe the expenditure of the poor. Also, the estimated poverty measures using the truncated gamma distribution are calculated.

REFERENCES

- [1] Aban, I.B, Meerschaert, M.M and Panorska, A.K. Parameter estimation for the truncated Pareto distribution, American Statistical Association, 2006, Vol. 101, No. 473, pp. 270--277.
- [2] Bandourian, R, McDonald, J.B and Turely, R.S. Income distributions: A comparison across countries and time, Discussion paper 2002, Brigham Young University, USA.
- [3] Berger, R.L and Sinclair, D.F. Testing hypotheses concerning unions of linear subspaces, Journal of the American Statistical Association 1984, vol. 79, pp. 158--163.
- [4] Chapman, D.G. Estimating the parameters of a truncated Gamma distribution, The Annals of Mathematical Statistics, 1956, vol. 27, No. 2, pp. 498--506.
- [5] Chang, C.H, Lin, J.L, Pan, N and Chiang, M.C. A note on improved approximation of the Binomial distribution by the Skew-Normal distribution, The American Statistician, 2008, vol. 62, No. 2, pp. 167--170.
- [6] Davidson, R and Duclos, J. Statistical inference for stochastic dominance and for the measurement of poverty and inequality, {it Econometrica}, 2000, vol. 68, pp. 1435--1464.
- [7] Duclos, J and Arrar, A. Poverty and Equity: Measurement, Policy and Estimation with DAD, Springer Science+Business Media, New York, 2006.

- [8] Egyptian Central Agency of Statistics. (1996). Income, Consumption and Expenditure Research (1995/1996), Egyptian Central Agency of Statistics, Cairo, Egypt.
- [9] Egyptian Central Agency of Statistics. (2000). Income, Consumption and Expenditure Research (1999/2000), Egyptian Central Agency of Statistics, Cairo, Egypt.
- [10] Egyptian Central Agency of Statistics. (2005). Income, Consumption and Expenditure Research (2004/2005), Egyptian Central Agency of Statistics, Cairo, Egypt.
- [11] Egyptian Central Agency of Statistics. (2009). Income, Consumption and Expenditure Research (2008/2009), Egyptian Central Agency of Statistics, Cairo, Egypt.
- [12] Foster, J.E, Greer, J and Thorbecke, E.A. A class of decomposable poverty measures, *Econometrica*, 1984, Vol. 52, pp. 761--766.
- [13] Gustafsson, B and Nivorozhkina, L. Relative poverty in two egalitarian societies: A comparison between Taganrog, Russia during the Soviet era and Sweden, *The review of Income and Wealth*, 1996, Series. 42, pp. 321--334.
- [14] Harrison, A. Earning by size: A tale of two distributions, *The Review of Economic Studies*, 1981, vol. 48, pp. 621--631.
- [15] Kakwani, N. Statistical inference in the measurement of poverty, *The Review of Economics and Statistics*, 1993, vol. 75, N. 4, pp. 632--639.
- [16] McDonald, J.B. Some generalized functions for the size distribution of income, A generalization of the beta distribution with applications, *Econometrica*, 1984, vol. 52, pp. 647--663.
- [17] McDonald, J.B and Xu, Y.J. A generalization of the beta distribution with applications, *Journal of Econometrics*, 1995, vol. 69, pp. 427--428.
- [18] Mudholkar, G.S and Huston, A.D. The epsilon-skew-normal distribution for analyzing near-normal data, *Journal of Statistical Planning and Inference*, 2000, vol. 83, pp. 291--309.
- [19] Oti, E.B, Kanbur, R and McKay, A. A poverty profile for Ghana, Working paper, 1990, World Bank, Washington. D.C. USA.
- [20] Salem, A.B.Z and Mount, T.D. A convenient descriptive model of income distribution: the gamma density, *Econometrica*, 1974, Vol. 42, pp. 1115--1127.
- [21] Sahn, D.E and Stifel, D.C. Poverty comparisons over time and across countries in Africa, *World Development*, 2000, Vol. 28, pp. 2123--2155.
- [22] Shorrocks, A.F. Notes and comments: Revisiting the Sen poverty index, *Econometrica*, 1995, vol. 63, pp. 1225--1230.
- [23] Sen, A.K. Poverty: An ordinal approach to measurement, *Econometrica*, 1976, Vol. 44, pp. 219--231.
- [24] Singh, S.K and Maddala, G.S. A function for size distribution of incomes, *Econometrica*, 1976, Vol. 44, pp. 963--970.
- [25] Shao, J. *Mathematical Statistics* (2nd ed), Springer Science+Business Media, New York, 2003.
- [26] Slocumb, J, Stauffer, B and Dickson, K.L. On fitting the truncated Lognormal distribution to soecies-abundance data using maximum likelihood estimation, *Ecological Society of America*, 1977, Vol. 58, No. 3, pp. 693--696.
- [27] Zheng, B. Statistical inference for poverty measures with relative poverty lines, *Journal of Econometrics*, 2001, vol. 101, pp. 337--356.